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Abstract

There has been a growing literature in both the US (for example Haurin and Brasington 1996, and Black 1999) and the UK (for example Gibbons & Machin, 2001) that estimates the way in which school quality is capitalised into house prices. Cheshire and Sheppard 1995 and 1999 estimated hedonic models in which the quality of the secondary school to which a household was assigned was a significant variable. This provided evidence that the value of secondary school quality was being capitalised into the price of houses.

In contrast Gibbons and Machin concluded that primary schools had an identifiable and significant price associated with their quality but that secondary schools did not. Their study did not have data for individual houses but used post-code sector data and then standardised for all but one variable: either the notional primary school catchment area or the notional secondary school catchment area.

Each of these analyses is predicated on the assumption that the value of local schools should be reflected in the value of houses. We expect variation in the capitalised price of a given school quality at either primary or secondary level according to the elasticity of supply of 'school quality' in the local market. This will vary systematically between and perhaps within cities and this paper explores the sources and the impact of such variations as well as the impact of model specification. The results support the conclusion that both secondary and primary school quality is capitalised into the market price of houses and that the capitalisation of school quality is discounted in areas where new construction is concentrated. We also find evidence that appropriate model specification is imperative since bias is evident both when key neighbourhood characteristics are omitted and if the actual allocation of addresses to schools is not included.

JEL: D12; H4; I2; R5

1. Introduction¹

Concern over the quality of local schools, and over the variation in this quality, has drawn the attention of parents, policy makers and scholars. For many households, there is a single path to access quality education: identify an acceptable quality state-supported school and purchase a house in the area served by that school. Households lacking the means to move to such areas will face reduced educational opportunities, and that fact continues to generate concern.

Interest in these issues has a long history. For economists, it goes back at least to Tiebout (1956) and Oates (1969). The questions they addressed were how do we determine the demand for and supply of local public goods, including education, and how do we pay for such goods. It was Oates who first drew attention to the ways in which the value of local public goods were capitalised in urban land markets. From this many implications flow including the role that land markets play in articulating social segregation (see for example, Brueckner, Thisse and Zenou, (1999)) and the interaction this will have with the distribution of incomes (see for example Cheshire, Monastiriotis and Sheppard, (2000)) and the supply characteristics of local public goods and amenities. In this paper we explore the extent of capitalisation of educational quality into house prices, and examine how this might be affected by land use planning policies.

At least three methodological approaches can be distinguished in the literature concerned with estimating the value placed on school quality. The longest established is a straightforward hedonic approach of which the other two are variants. The hedonic approach has some 80 years of evolutionary development behind it since agricultural economists first implemented it as a purely empirical technique to help identify the characteristics of vegetables commanding the highest price. Since Rosen's (1974) contribution it has become one of the standard techniques for analysing the price of complex goods, particularly that of housing.

Over the past 25 years a great many new insights have been gained particularly as to the importance of model specification and the way in which the values of local neighbourhood characteristics, local public goods and locationally specific amenities are capitalised into land values. In parallel there have been important technical innovations in the effort to capture these effects more precisely. Perhaps the single most important lesson that has been learned is the most obvious: the value of any house varies

¹ We would like to thank the Leverhulme Foundation and the Lincoln Institute of Land Policy for supporting the work underlying this paper.

systematically and substantially with its location and these location-specific factors are at least as important as the characteristics of the structure itself in determining market price.

Because the relationship between market price and characteristics is typically non-linear, the specification of hedonic models is critical in determining the prices estimated for individual characteristics. Poorly specified models can yield misleading results. For example, the values of omitted locationally specific characteristics tend to be attributed to the estimated price of space, either internal space in the house or land area. Most of the value in the market price of urban land is in fact represented by the capitalised value of locationally specific goods. These include the quality of local schools.

This may underlie the concerns that have led researchers recently to search for other ways of isolating the values attached to particular local public goods (or other spatially determined amenities). Both the methods deployed in the recent literature on the value of schools are essentially variants of hedonic analysis. Black (1999) sought to isolate the value placed on school quality by taking a large sample of house values for which she could reasonably argue that the only difference between them was the quality of the schools to which they gave access. In so far as this was correct then it followed that one could attribute differences in their value to differences in school quality.

This 'matched pair' method is really a type of hedonic analysis. It is implicitly admitted that many variables or attributes determine the price paid for the complex good housing and the researcher is simply trying to set up a situation in which the influence of all but one is eliminated. A difficulty with the approach is that there are no obvious tests to apply to see to what extent the research design has succeeded. In so far as there are omitted spatially fixed effects that are correlated with the school districts then there would be bias in the estimated value assigned to schools. In a fully specified hedonic model however we can (and should) undertake tests of model specification.

Gibbons and Machin (2001) develop another variant on hedonic analysis. They employ a kernel-based technique to offset for spatial fixed effects and exploit the co-variation in house prices and school performance within narrowly defined spatial units to reduce the need for a large set of covariates. They use mean house prices by area and deviations from means. There are some potential problems with this approach. One relates to the characteristics of supply which, as is discussed below, will vary from city to city and under some circumstances, will vary systematically by location within cities reflecting the quite local elasticity of supply of housing (the implications of which are explored by Hilber and Mayer 2001). Thus the resulting estimates will be, at best, mean values for the whole area

analysed (in the case of Gibbons and Machin, England and Wales) and may conceal very large variation between areas. Indeed it is perfectly possible that in some areas primary school quality is more expensive whereas in others, secondary school quality is more expensive.

A second problem with this approach is really the same as the criticism of Black's matched pairs approach made above. While one may design the technique to control for spatially fixed effects – such as neighbourhood characteristics, other local public goods and specific locationally fixed amenities – we cannot test for the extent to which one has succeeded. Some of these locationally fixed effects are very local (for example views, access to local amenities, local disamenities from industrial land use, noise disturbance or the socio-economic characteristics of the neighbourhood). Since the catchment areas of primary schools are small, failure to separately account for spatially fixed effects will tend to be reflected in the value of the estimated parameter for primary school quality. Sorting processes in housing markets concentrate socio-economic groups whose children do better in the educational system in precisely the same areas, exacerbating the upward bias to the estimated value of primary school quality.

For these reasons we use a traditional hedonic approach and attempt to measure a wide range of local neighbourhood characteristics, including the socio-economic composition of the neighbourhood and other local public goods and localised amenities. We have also included the most fundamental of all features of the structure of urban land markets – land consumption and accessibility to jobs.

2. The supply of quality and capitalisation into house values

The economic and institutional structure within which educational opportunities are made available will naturally influence our ability to measure the willingness to pay for school quality by affecting the extent to which this valuation is reflected in house prices. While the demand for school quality may not vary greatly from one city to another, at least within the same country, the implicit price may vary because of variation in the supply of school quality available to a household living at a specific address. This can vary substantially from one city to another. Where educational opportunities are at least in part determined by residential location, there are three central factors that are relevant in determining the supply of quality and the extent of capitalisation. These are the elasticity of housing supply, the availability of substitute providers of education (other than the state-supported provider designated for the particular address), and the anticipated risk of variation in the quality of education provided. We discuss each of these factors in turn.

A. Housing Supply

An important source of variation in educational opportunity is the availability of housing in the areas served by (better) quality schools. The responsiveness in housing supply, in turn, is determined by construction costs and local planning regulations. Cross sectional variance in the elasticity of housing supply is largely determined by variation in planning regulations. If in one location the supply of houses is fixed whilst it is highly elastic in another, then the measured capitalisation of school quality will vary even though demand is invariant. This implies that the possibility of local variation in the implicit price of school quality cannot be entirely discounted.

Variation between cities in land use regulation implies that we may observe substantial differences in the supply characteristics of school quality between cities. Furthermore, there may be differences in the supply of school quality within cities. This may arise because of differing elasticities of supply of housing according to location. Cheshire and Sheppard (1995) identified substantial differences in the degree of planning restriction on housing supply between cities that corresponded with differences in the capitalised price of secondary schools. The market from which the data for the present study are drawn is subject to restrictive planning controls, so there will be a relatively inelastic supply of housing in the whole area but localised housing supply will vary from location to location within it as particular parcels of land are released.

The impact of local variation in housing supply elasticity, and its impact on capitalisation has been the focus of recent research in US housing markets. Hilber and Mayer (2001) and Brasington (2002) have drawn attention to the fact that the extent of capitalisation may be reduced in areas where housing supply would be expected to be more elastic. Comparing across cities in Massachusetts, Hilber and Mayer find empirical support for the observation. Comparing central with peripheral residential properties in urban areas of Ohio, Brasington finds that the capitalised value of a given level of quality is reduced for houses at the edge of the urban area.

It is possible that observed reductions in capitalisation might exist for other reasons, related to the availability of substitute sources of education or variance in quality discussed below. Before proceeding to consider these factors, we note a final explanation related to the regulation of housing supply. An apparent discount in the implicit price of school quality could be due to a land use planning system that concentrates new construction in localities with significant local disamenities (and hence reduced opposition to new development), where the disamenities are difficult to measure and control

for in the hedonic model. In this case apparent discounting of school quality might reflect the impact of such omitted variables.

B. State School Quality and the Availability of Substitutes

Even if the supply of housing were completely inelastic, house values would be little affected by school quality if substitute sources of quality education were readily available. In such cases, the impact of state-supported school quality on house values would be an accessibility premium related to proximity to the school. For primary schools in particular this 'distance decay' might still be significant. The cost of sending a child to a more distant primary school is substantially higher than sending a child to a more distant secondary school. Children younger than 10 or 11 will normally be taken to school by a parent. Secondary school children will usually not require parental effort for the journey to school. Thus there may be a distance cost associated with primary school quality although its magnitude will depend on the size of the catchment areas within which pupils from the school live.

In the city from which our data are drawn, there are four possible substitutes for the dedicated state school (and these may vary between the primary and secondary levels). These are (1) a private school, (2) a parochial (church-affiliated) school, (3) to gain admission to a state-supported 'Grammar School' and (4) to request transfer to a state-supported school other than the one identified to serve the address.

The market we study is a relatively high-income community, well endowed with private schools, particularly at secondary level. This suggests there will be an upper limit on the capitalised price of school quality. Access to private schooling is controlled largely by income not location, so if a given degree of school quality can always be purchased in the private market for educational services, this price will determine the upper limit of the capitalised value of state school quality, but the cost of private schools implies that this upper limit will be relatively high.

At the primary school level (i.e. for children below the age of 11) there are a variety of state-funded parochial schools, admission to which is more loosely related to home address. While in some neighbourhoods this might be a factor, there are indirect costs associated with parochial education. For example, Gibbons and Machin (2001) point out that such education entails a cost associated with conforming to religious requirements. This suggests that parochial schools may be similar to the availability of private education in its overall impact: it places an upper limit on the estimated value of educational quality, but this limit will tend to be relatively high.

An unusual feature of the school system in our sample area is the continued existence of Grammar Schools (state-supported secondary schools with entry highly selective according to tested academic ability). It is possible that this means that as well as a price cut-off determined by the price of private schooling, there is also a quality cut-off in terms of capitalised values. Parents of very high ability children who expect them to get into the local Grammar School might be unwilling to pay for school quality by moving to the catchment area of a better, non-selective school, since they expect their child to get into the Grammar School. There might thus be an apparent drop off of school quality-price at the highest level of measured output quality.

The final factor that might limit the extent of capitalisation of school quality into house prices is the possibility that a household requests and is granted permission to send their children to a state-supported school other than the one to which the house would usually be allocated. The frequency with which such requests are successfully made determines the overall 'porosity' of school catchment zones, and this will be determined by a combination of factors.

In England (as in many other countries) each house is assigned to a default primary and secondary school. Parents may in principle nominate any school for their child but presumably there is considerable inertia: most parents simply accept the local school. In requesting a different school to the local one parents are presumably guided by their perception of the probability of such a nomination being successful. If they choose a school other than their default school and the local education authority (LEA) do not accept this choice then parents may appeal. Again it is likely that in deciding whether to appeal parents take some account of the probability of success since the appeal process takes some time and effort.

In trying to compare differences in the underlying 'porosity' – the probability that a child living at a particular address will actually attend the local default school in the catchment area of which the house is located – we are hampered by only being able to observe some of the relevant variables. There are no data on the proportion of parents choosing a non-local school nor on the proportion of such choices that are rejected by the LEA. Since 1997, however, there are systematic data on the appeals process. Figures are published for all LEAs in England on the total number of admissions to primary and secondary schools, on the number of appeals against the allocations made by parents and the outcome of these appeals. Some of these data are summarised in Table 1.

Authority	Primary Schools							Secondary Schools									
	1997-98		1998-99		1999-00		M	Mean		1997-98		1998-99		1999-00		Mean	
	Succ ¹	Flex ²															
England	1.7	31.0	1.6	29.0	1.3	25.4	1.5	28.5	1.8	23.3	2.0	23.5	2.3	23.5	2.0	23.4	
Reading Area ³	1.6	31.4	3.7	31.2	1.7	17.3	2.3	26.6	0.4	7.5	0.6	15.6	0.5	9.6	0.5	10.9	
Inner London	1.5	16.0	1.2	15.1	0.6	8.9	1.1	13.3	1.8	15.0	1.7	9.3	1.6	8.3	2.1	10.9	
Gtr. London	2.7	23.6	2.1	20.3	1.3	15.7	2.0	19.9	2.2	13.9	2.2	11.9	2.7	13.3	2.4	13.0	
Oxfordshire	1.3	43.0	1.6	47.6	1.2	42.4	1.4	44.3	1.5	38.4	2.1	45.4	1.5	33.3	1.7	39.0	
Darlington	7.1	44.9	8.8	40.7	6.0	41.5	7.3	42.4	4.1	34.3	4.7	37.1	2.5	31.6	3.8	34.3	
Nottingham	0.4	30.3	0.8	28.4	1.4	35.2	0.9	31.3	0.5	21.7	0.8	17.6	1.0	23.7	0.8	21.0	
Manchester	2.2	28.0	2.8	38.7	1.7	25.7	2.2	30.8	1.2	14.7	2.1	24.1	2.1	18.2	1.8	19.0	
Cheshire	0.1	27.5	0.7	69.2	0.7	51.3	0.5	49.3	0.5	62.7	2.0	59.3	3.9	65.0	2.1	62.3	

Table 1: Success rate of Appeals against School Allocation and per Appeal

¹Successful Appeals as % of Total Admissions ²Successful Appeals as % of Total Appeals ³Weighted mean for three Local Education Authorities: weights determined by distribution of sampled houses

Patterns are reasonably consistent between LEAs. Darlington has a high rate of successful appeals, and Cheshire has a low rate relative to admissions but a high proportion of the appeals that there are, are successful (perhaps indicating a very flexible policy with a high rate of unobserved nominations of non-local schools as well). Inner London has a low rate of both successful appeals relative to admissions and relative to appeals. This may reasonably be interpreted as indicating an inflexible regime in which the school a child attends is largely determined by home address.

The data for our sample area (Reading) suggest a regime that is rather less restrictive that that of Inner London at the primary level but even more restrictive at the secondary level. Looking at the mean rates for the three years only 0.5 percent of children successfully appeal against their secondary school allocation in the Reading area (one quarter the success rate of Inner London or England as a whole) and 10 percent of appeals are successful – the same as Inner London but half the proportion of England and one sixth that of Cheshire.

C. Risk of Variation in School Quality

Since LEAs may establish catchment area boundaries, it is expected that boundaries will be revised on a regular basis in order to fill available school places and eliminate spare capacity. Officials from the LEAs responsible have confirmed that this is standard practice in our sample area. This creates uncertainty for the buyer concerning which school will serve the house in the future. At the urban periphery where population density is lower, the spatial magnitude of such adjustments would be expected to be larger, and thus the uncertainty concerning distance to school is larger for these properties. Furthermore, since the limited land release for development that does occur takes place mostly at the urban periphery, there is for such properties an added source of uncertainty regarding the nature of the population who will be served by the school. The combined impact of these

considerations is to impose greater risk of variation in school quality in some areas (primarily the periphery and, particularly in areas of new construction) of the market than in others.

The evidence reviewed above supports the conclusion that in the area from which our sample is drawn, most children go to the school determined by the location of their home and by the boundaries of the catchment areas in force in the year they first go to either primary (at 5 years) or secondary school (at 11). This probability is significantly higher at the secondary school level and is very high by the standards of England as a whole.

Let us summarise our observations concerning the elasticity of supply of school quality (or at least supply as measured by those variables we are using to capture it). If parents are concerned to increase the probability of their child(ren) attaining a particular level of qualification then their choice at the primary level will be between: a secular state primary, a parochial school or a private school. If they choose a state school then they can move to the catchment area of the school of their choice, trading off price against quality; or they can try to obtain entry to a more distant school, probably a parochial one, and pay a price in the journey to school and church. *De facto* there is more flexibility (that is an ability to exercise choice of school) at primary than at secondary school level. These considerations suggest the supply of school quality at primary level may be more elastic than it is at the secondary school level.

At the secondary level parents can make similar choices except that in Reading there is a strong constraint against choosing any secondary school other than the one in the catchment area of which they live. Boundaries of catchment areas are revised annually, adding some uncertainty to the correspondence between school and residential address and this uncertainty is likely to be systematically higher in lower density areas on the periphery and in those areas where new construction is concentrated as LEAs attempt to manipulate intakes to utilise school capacity.

How are these restrictions on choice affected by the land use planning regulations? Development controls in Reading effectively impose a non-price constraint on housing supply and reduce the elasticity of supply. In the most extreme case, the effective supply of houses in any school's catchment area would be given by the existing stock. In this situation, we would expect to see greater capitalisation of the value of educational quality into house prices.

3. Data and Setting

Our data are drawn from the urban area of Reading, England. The city is located on the Thames about 35 miles west of central London. Reading is subject to considerable pressure for growth and residential development, and has in response adopted some of the most restrictive planning policies in England and Wales. With frequent high-speed rail links to London, proximity to Heathrow airport and other locational advantages the area has attracted a number of high technology firms² and more generally follows the development patterns typical of prosperous, middle-size cities of the southeast of England. Despite its proximity to London, Reading is a major employment centre with more than 85 percent of its employed residents working locally and a strong central business district employment concentration. It is a reasonable city, therefore, to which to apply the familiar monocentric model of urban land use.

In 1991 the city had a metro area population of approximately 337,000 persons comprising 129,000 households. At the time of the 1999/2000 survey we estimate that there were 131,370 households. Our initial sample of properties comprised over 870 separate structures. This provided a sample of approximately 20% of the residential properties offered for sale by major estate agents during the 17 months covered by the data. Complete data including location, structure characteristics, sales date and price, and school assignments were available for 490 observations and these are used in the analysis below.

Supplemental information on land use was assembled from Ordnance Survey resources and aerial photographs. Data on both secondary and primary school catchment areas was obtained from the local education authorities. Data on state-supported school quality were obtained from the Department of Education website³. The measure used for primary schools was the performance of its pupils on the Key Stage 2 tests⁴. For secondary schools the measure of school quality was the proportion of pupils obtaining 5 or more passes at grade C or better in GCSE⁵. Data on the availability, performance and price of local private schools was obtained from the ISIS website. The Department of Local Government, Transport and the Regions' (DETR) index of employment deprivation was used as the measure of the socio-economic characteristics of the neighbourhood. Appendix Table 1 provides some descriptive statistics for the sample and a description of each variable used in the analysis.

² Microsoft, Oracle, Hewlett-Packard and others

³ http://www.dfes.gov.uk/statistics/DB/SBU/b0333/index.html

⁴ Tests administered nationwide and designed to assess achievement in mastering that portion of the national curriculum, known as 'Key Stage 2', deemed appropriate for ages 7 to 11.

⁵ A nationwide exam taken at minimum school leaving age, 16.

4. The Hedonic Model

Our basic model follows a procedure similar to that used in Cheshire and Sheppard (1995, 1998). We locate each house in the sample and measure the size of the plot of land associated with it. We then estimate a modified linear Box-Cox hedonic price function given in equation (1). Note that the value function for urban residential land, specified in equation (2), is estimated directly as part of the hedonic price function. The land rent is 'monotonic' only in the sense that it is radially symmetric so that land value must increase or decrease at the same rate in any given direction away from the urban centre.

$$\frac{\mathbf{P}^{\Psi} - 1}{\Psi} = \mathbf{K} + \sum_{i \in D} \beta_i \cdot \mathbf{q}_i + \sum_{j \in C} \beta_j \cdot \left(\frac{\mathbf{q}_j^{\lambda_1} - 1}{\lambda_1}\right) + \sum_{k \in E} \beta_k \cdot \left(\frac{\mathbf{q}_k^{\lambda_2} - 1}{\lambda_2}\right) + r(\mathbf{x}, \theta) \frac{\mathbf{L}^{\xi} - 1}{\xi}$$
(1)

where:

Р	=	sales price of structure
q_{i} , q_{j}	=	structure and location-specific characteristics
K, β_i , λ_i , ψ , ξ	=	parameters to be estimated
L	=	quantity of land included with structure
D	=	set of indices of characteristics which are dichotomous
C	=	set of indices of characteristics which are continuously variable
E	=	set of indices of characteristics measuring educational quality
$r(x, \theta)$	=	land rent function defined below
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and the land rent function is given by:

$$\mathbf{r}(\mathbf{x},\boldsymbol{\theta}) = \beta_1 \cdot \mathbf{e}^{\mathbf{x} \cdot (\beta_2 + \beta_3 \cdot \mathbf{sin}(\mathbf{n} \cdot \boldsymbol{\theta} - \beta_4))}$$
(2)

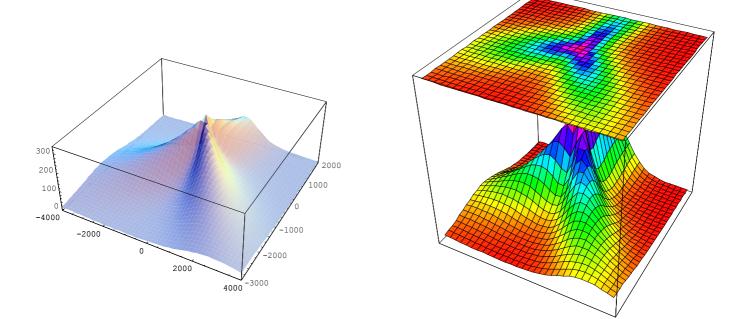
where:

x=Distance from the city centre θ =Angle of deflection from the city centren=Number of 'ridges' in land value, representing radial asymmetries β_i =Estimated parameters of land value function

Searching over a small grid (1-4) it was determined that a rent function with n=3 ridges provided the best fit to the data. The estimated land value depends on the location and also the size of the plot and

type of structure built upon it. For a structure matching the sample mean in all attributes (except location) the spatial structure of the land value function is illustrated below in Figures 1 and 2.

The surface is viewed from the southeast looking towards the northwest. The three ridges closely track the local transport system. They are aligned with the main road access routes to the city centre: the A329M linking the main London Bristol motorway – the M4 – to the centre from its eastern junction; the access route from the M4 at its junction to the south of the city along the A33; and the main route, again linking to the M4, to the west of the city along the A4.



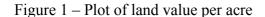


Figure 2 – Land value with projected contours

The measure of the value of land shown in Figures 1 and 2 is essentially the price of 'land as pure space with accessibility'. Actual market prices of vacant land include the capitalised value of all the local amenities, neighbourhood characteristics and local public goods to which occupation of the land gives access. As was shown in Cheshire and Sheppard (1998) these amenity values may exceed the value of land as pure space with accessibility⁶.

5. Interpreting the results

We present the results of estimation of six models. Parameter estimates for the reported models are presented in Appendix Table 2. Model I presents estimates of a basic model including measures of the quality of the primary and secondary schools to which the address is assigned by the local education authority. Model II presents an estimate of the same model, but using the quality measures of the

⁶ In the data studied in Cheshire and Sheppard (1998) the amenity values were greater by a factor of up to eight

primary and secondary school that are nearest (straight-line distance) to the house. Model III presents a model using the measures of school quality at the assigned schools, but drops the DETR Employment Deprivation index⁷, and Model IV repeats this structure using the school quality measures from the nearest schools.

The last two models include all available variables plus an index for the house being located in an area of the urban periphery that has experienced considerable new construction. In Model V this index is included in a way that allows estimation of any discounting of the value of school quality for houses in these areas; in Model VI a simple dummy variable is incorporated if the house is located in a (peripheral) area within which new construction has been concentrated.

In all of the models we use the reported educational quality that would have been the 'most recent' at the time the house was sold. There is variance from year to year in exam scores, so that one might be concerned that an average of recent scores might be a more appropriate measure. We have tested use of either a 5-year average or 3-year moving average for exam scores. In neither case was there consistent improvement in model performance, and for the majority of cases the model performance was worse. For this reason we use the 'contemporaneous' exam results⁸.

A. Value of Primary and Secondary Schools

We start by addressing the question: which types of schools are of greater value to purchasers of houses? There are at least two different approaches to this question, and it turns out (at least with the data sample used for this analysis) that each approach gives a somewhat different answer.

The first approach is simply to compare the estimated hedonic prices of each measure of school quality. Examination of the parameter estimates in Appendix Table 2 shows immediately that the estimated parameter for the quality of secondary schools is considerably larger than for primary schools (as well as having a larger t value associated with it). A better comparison is afforded if we standardize the ranges of the quality measures. Figure 3 presents plots of the hedonic price (in thousands of pounds⁹) for both the measure of secondary school and primary school quality, standardized so that the movement from 0 to 1 represents the total possible range of outcomes in the quality measure. At comparable levels, the secondary school quality is 'more valuable'. It is notable

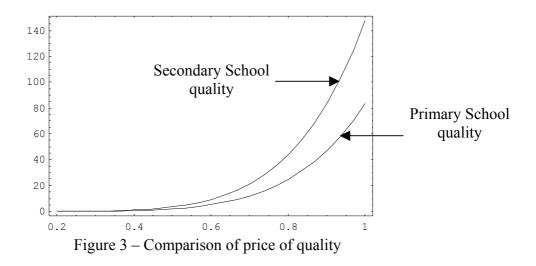
⁷ Various of the available deprivation indices were tried. The multiple index of deprivation worked best in a statistical sense but, because one small element of that is the performance of the local primary school on Key Stage 2, the results obtained using the employment deprivation index are shown here.

⁸ The results using averaged exam scores are available from the authors.

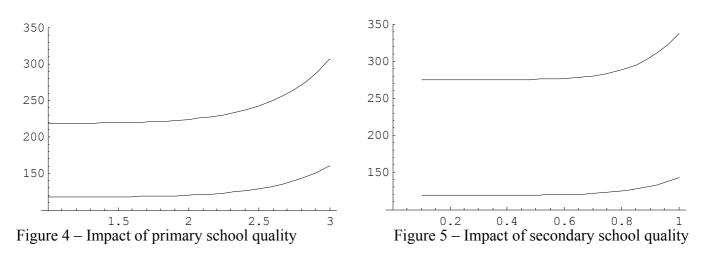
⁹ Evaluated for a house whose value and other characteristics are equal to sample mean values.

how non-linear the price paid for school quality appears to be; better quality really only commands a substantial price in the top one third of the school quality distribution.

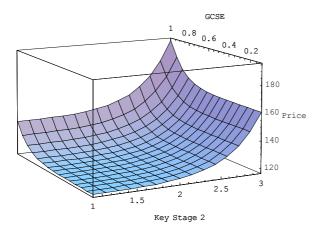
An alternative approach is to ask which factor contributes the most to the value of houses within our sample. This question is different because of differences in the range of measured school qualities. The movement from the 'best' to the 'worst' secondary school within the area may be a much different proposition than the movement from the 'best' to the 'worst' primary school. Indeed, this is confirmed by considering the change in value of an average house as we move from the lowest to highest measured quality in the sample.



Figures 4 and 5 provide one way of examining this issue. Each figure shows the variation in the predicted value of the average (bottom curve) and the most expensive (top curve) house in the sample as school quality varies from the lowest observed level to the maximum possible (the vertical axis in both cases is measured in thousands of pounds).



Figures 6 and 7 provide a visual representation of the joint impact of school quality of both types on the price of an 'average' house, along with the distribution of observations in the sample within different ranges of the school quality spectrum. Figure 6 provides a surface that illustrates the impact on house values of changes in both primary (Key stage 2) and secondary (GCSE) school quality. Figure 7 superimposes this surface over a 'histogram' that shows the share of sample observations within each range of qualities. It is apparent that the distribution of state-sector secondary schools is concentrated in the lower to middle quality range, while the distribution of state supported primary schools covers a broader range of school quality levels.



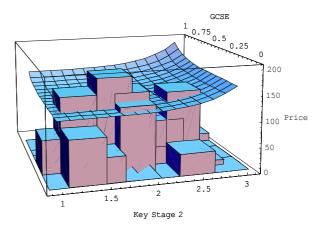


Figure 6 – Impact of quality on house price

Figure 7 – Impact and distribution of quality on price

In summary, while the hedonic price of secondary school quality is higher than the price of primary school quality, moving from the worst to the best possible secondary school would increase the value of the average house by £23,763 (or 18.7 percent of the value of a mean house). Moving from the worst to the best possible primary school would increase the value of the average house by £42,541 (33.5 percent of the mean house value). In passing it may be noted that the estimated value added to the price of a mean characteristics house moving it from the catchment area of the worst to the best secondary school from the 1993 sample was an increase of 14.1 percent (Cheshire, Monastiriotis and Sheppard (2000)). The hedonic model used for the 1993 data did not include primary school quality since Key Stage 2 test results were not available then.

B. Models With Measurement Error: Nearest vs. Assigned Schools

Assignment of houses to a default primary and secondary school is up to the LEA. This information is not available from any central source, and for some education authorities can be difficult to obtain. For

this reason many studies of the effects of school quality do not actually use the quality level of the assigned school, but rather the quality level of the school (primary or secondary) that is located nearest to the house. While this is feasible, there is a question as to whether this provides a good approximation of the school quality that would actually be available to the residents of a particular house.

Appendix Table 3 presents the correlations (across properties in our sample) between quality variables for assigned and for nearest schools. It is immediately apparent that the correlation between the quality measures for assigned schools and closest schools is low. In the case of secondary schools, which have larger catchment areas, the R is only 0.435. Comparing the estimated parameters for models I and II (shown in Appendix Table 2) we see that using quality measures for schools actually assigned to addresses provides a better fit for the data than using the values for the closest school. The t values for the relevant parameters fall and that for primary schools ceases to be significantly different from zero.

These results suggest that caution is certainly appropriate when interpreting estimates based on measurement of school quality using the nearest school rather than the assigned school. Within the sample underlying the present analysis, the nearest school is at best a weak approximation of the school quality actually available to a child who attends the school designated by the LEA.

C. Quality of Schools and Neighbourhoods: Estimation With Omitted Variables

A further concern in the evaluation of school quality arises because the school catchment area, particularly for primary schools, may serve as an approximation for local neighbourhood effects. Therefore omitted variables, particularly those related to neighbourhood quality, may bias the estimates of the value of educational quality. Such estimates will reflect both the value of education and the value of the omitted neighbourhood variables. To examine this issue we examine the effect on model estimation when the DETR employment deprivation index variable is dropped. This variable provides a measure of concentration in the neighbourhood (census ward) of persons having little success in the local labour market. It therefore helps to capture the socio-economic character of the neighbourhood.

Dropping the measure of the socio-economic character of the neighbourhood substantially increases the estimated value of the primary school parameter – it increases sevenfold in absolute terms – but because it impairs the overall performance of the model its t value is still lower than in Model I. The estimate for the secondary school parameter falls in absolute terms if the deprivation index is dropped,

although it remains statistically significant. This provides justification for concern that there is likely to be an upward bias in the estimated impact of primary school quality on house prices if other important local neighbourhood effects are not independently controlled for.

D. Discounting at the Urban Periphery: Planning for Growth or Uncertainty?

Finally, we turn attention to the possibility that the value of educational quality is not fully capitalised into houses located at the urban periphery in areas where new construction is concentrated. As discussed above, there are three potential reasons why this might occur. In a city subject to land use regulation there will be greater elasticity in the supply of developable land in those areas where land release occurs. This implies that some of the increased demand for housing is accommodated by increases in supply, so the price will not rise by the full increase in consumer willingness to pay for a house with access to high quality schools¹⁰. If this generates a different level of capitalisation it implies disequilibrium in the housing market, since residents are able to 'buy' a given level of school quality at lower cost in the areas of new construction than they could elsewhere¹¹.

A second possibility is that the planning system operates in such a way as to concentrate new development in localities with disamenties of some sort. Development in such areas generates less opposition from local residents. If these disamenities are not captured in the hedonic then the specification error appears as a reduced level of capitalisation.

A third explanation arises due to the uncertainty regarding school quality in rapidly growing areas. This uncertainty arises from two sources. First, school quality is sensitive to both the quantity and quality of student intake. Both of these may exhibit considerable variance in peripheral areas. Therefore, house buyers may be uncertain as to the exact quality of schools that will be available to them. They therefore discount the amount they are willing to pay for current school quality to reflect this risk. Furthermore, the designation of school catchment areas will be subject to greater and more frequent change as the LEA seeks to equalise school intakes and this adds uncertainty both as to school quality and the distance to be travelled to school.

To see if such discounting appears to be present in Reading, we estimate two modified hedonic models. One of the models, reported as Model V in Appendix Table 2, has the form:

¹⁰ This is the argument made in Brasington (2002).

¹¹ Any difference in commuting costs is already compensated by reduced land prices at the periphery.

$$\frac{P^{\psi}-1}{\psi} = K + \sum_{i \in D} \beta_i \cdot q_i + \sum_{j \in C} \beta_j \cdot \left(\frac{q_j^{\lambda_1}-1}{\lambda_1}\right) + \sum_{k \in E} \beta_k \cdot \left(1 - \beta_P \cdot \delta\right) \cdot \left(\frac{q_k^{\lambda_2}-1}{\lambda_2}\right) + r(x,\theta) \frac{L^{\xi}-1}{\xi}$$
(3)

where all variables are as defined above, and in addition:

 β_P = estimated parameter to capture the reduced impact of educational quality at the periphery

 δ = dichotomous variable taking the value 1 for houses located in peripheral areas of new construction, and 0 otherwise

The second, Model VI, simply uses the dichotomous variable δ as a separate characteristic. This allows us to test whether any discount strictly relates to school quality or just reflects unmeasured negative effects (disamenities) in such areas.

Estimates of Model V show that there is indeed a strong discounting of school quality in wards where new construction was concentrated. Since this model clearly outperforms Model VI we can safely conclude that the discounting relates strictly to school quality not to the areas' amenity levels. This discounting is reflected in the variable $\beta_{Periphery}$, whose value indicates that for houses located in the peripheral areas of new construction the value of educational quality is discounted by more than 60% relative to houses in other portions of the city. This is a very large discount, and given the limited new construction allowed by planning policies is unlikely to be entirely due to supply response. The two remaining possible explanations are not mutually exclusive, however; both could contribute to the discount observed. Further research is required to isolate the separate contributions of each factor.

6. Conclusions

In this paper we have sought to show that while average measures of the price of school quality estimated over many communities may be useful, because of local variation in the supply of school quality, one should expect that there would be substantial variation in the capitalised value of school quality between and even within cities. In addition we have highlighted what we see as the need to have as completely specified an hedonic model as possible if one is to obtain accurate measures of the capitalised value of school quality.

Applying such an approach to the city of Reading in South East England for data relating to 1999/2000 we find that the quality of both local secondary and primary schools was capitalised into house prices. The statistical significance of secondary schools was considerably greater as was the relative price that secondary school quality commanded. However there are far more primary schools

and the range in their performance is considerably greater. Thus there was a larger total impact on house prices associated with 'moving' a standard house from the worst to best primary school catchment area than there was in the case of a similar move between secondary school catchment areas. The price paid for school quality was substantial and, in the case of secondary schools for which a direct comparison is possible, comparable to estimates for 1993 in the same housing market.

Three further conclusions emerged from this analysis. The first was the need to include (at least in markets where school catchment areas are non-porous) the quality for the actual school serving the catchment area rather than the quality associated with the nearest school. Indeed there was only a low correlation between the quality measures for the two. The second is the danger of obtaining an upwardly biased measure of primary school quality if as full a range of local neighbourhood characteristics and amenities as possible is not included. Simply omitting the employment deprivation index for the local ward from the model increased the absolute value of the parameter estimate for primary school quality sevenfold (while reducing that of secondary schools).

The third conclusion is that school quality appears to be significantly discounted in areas in which new construction is concentrated. While this finding is consistent with the hypotheses of Hilber and Mayer (2001) and Brasington (2002) that the elasticity of supply of housing will influence the extent to which school quality is reflected in house prices, both our findings and theirs are capable of other explanations. The discount might reflect uncertainty as to future changes in school catchment areas in such neighbourhoods or uncertainty as to what school quality will actually be since catchment areas and intakes are subject to greater change. Our results do show, however, that it is unlikely that the discount reflects the influence of omitted local disamenities from the model since it attaches strictly to school quality rather than to the area itself. The two plausible explanations are not mutually exclusive – both could be true. An important goal of continuing research is to find techniques of distinguishing them.

In more general terms the results reported here confirm previous findings that access to better schools, whether these are provided free from taxation or through the market, is still conditioned on income. Poorer households face the same income constraint on access except that it operates through the market in housing rather than that in education.

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Variable	Mean	σ	Min	Мах	Description
Price	126.9378	48.6852	45	385	
Detached	0.0984	0.2981	0	1	1 if property is a detached house
Semi-detached	0.1687	0.3748	0	1	1 if property is a semi-detached house
Terrace	0.3896	0.4881	0	1	1 if property is a terrace house
Townhouse	0.1024	0.3035	0	1	1 if property is a townhouse
Parking	0.3153	0.4651	0	1	1 if property has off-street parking
Thames	0.0080	0.0894	0	1	1 if centre of lot is within 150 m of Thames
Rail	0.1104	0.3138	0	1	1 if centre of lot is within 200 m of rail line
Cul-de-sac	0.2209	0.4153	0	1	1 if property is located on a cul- de-sac
Minor Road	0.6386	0.4809	0	1	1 if property is located on minor through street
B-Road	0.0161	0.1258	0	1	1 if located on "B" class roadway
A-Road	0.0482	0.2144	0	1	1 if located on "A" class roadway
Time Trend	0.9351	0.3020	0	1.4740	Years since 6/1999 (time trend)
Bedrooms	2.5815	0.8436	0	6	Number of bedrooms
Baths	1.3448	0.6576	0	5	Number of bathrooms
Nosquare	0.6103	0.1814	0.1854	1.0408	Ratio of lotsize to perimeter
SqFt	676.1154	242.1323	189.8611	1749.0139	Square feet of internal living space in house
Industry	10.6827	11.7065	0	50	Percent of land within 1 km square in industrial use
EmployDepriv	7.0933	2.2435	2.4418	10.2846	DETR index of employment deprivation
Lotsize	222.6534	214.7078	22.10882	2054.5471	Lotsize in square metres
Distance	2289.1982	1462.9522	54.65398	3331.3380	Distance from town centre in metres
θ	-0.4863	2.0548	-3.1391	3.1391	Direction in radians from town centre (East=0)
PrimarySchool	1.8654	0.4713	1.14	2.84	Sum of share of pupils in assigned school passing keystage 2 exams in English, Math, and Science
SecondarySchool	0.3469	0.1390	0.1	0.75	Share of pupils in assigned school receiving a grade of C or better in 5 or more GCSE exams
PrimarySchool	1.8457	0.4650	1.14		Sum of share of pupils in nearest school passing keystage 2 exams in English, Math, and Science (Models II and IV)
SecondarySchool	0.3633	0.1356	0.05	0.72	Share of pupils in nearest school receiving a grade of C or better in 5 or more GCSE subject exams (Models II and IV)
Periphery	0.0944	0.2926	0	1	1 if Property located in peripheral ward with new construction

Appendix Table 1: Variable Descriptions and Descriptive Statistics

Parameter	Model I	Model II	Model III		Model V	Model VI
	3.124553					3. 046134
ρ ₀ t	29.601	3.230343 11.183				8.763
β _{Detached}	0.185303					0.206525
PDetached	8.469	4.751	12.608			3.551
$\beta_{Semi-detached}$	0.119002					0.133012
t	6.800	4.798				3.062
β _{Terrace}	0.051215	0.054739				0.056112
t	4.195					2.705
$\beta_{Townhouse}$	0.07224	0.080601	0.084975	0.081427	0.077399	0.080087
t	4.868	4.201	5.853	4.648	3.480	2.951
$\beta_{Parking}$	0.011386	0.010764	0.01165	0.007322	0.011901	0.012787
t	1.742	1.607	1.762	1.062	1.723	1.398
β_{Thames}	0.074639	0.091209	0.09254	0.107257	0.077876	0.080634
t	2.339	2.487	2.991	2.825	2.453	1.775
β_{Rail}	-0.00837	-0.00985	-0.0076	-0.00957	-0.00899	-0.007828
t	-0.949	-1.052				-0.645
$\beta_{Cul-de-sac}$	0.030018	0.03431	0.05332	0.050991	0.034378	0.035039
t	2.265	2.234				1.821
$\beta_{Minor Rd.}$	0.005123	0.006463	0.019676	0.018172	0.008648	0.008621
t	0.452					0.580
β_{B-Road}	0.099615	0.109639				0.112372
t	3.814					2.681
β_{A-Road}	-0.0013	-0.00385				-0.000335
t	-0.071	-0.227				-0.0149
β _{TimeTrend}	0.029917	0.034401	0.037374			0.034544
t O	3.185	2.931	4.153			2.526
β _{Bedrooms}	0.02032	0.024127	0.027871	0.025885		0.020019
t R	3.031 0.051009	2.955 0.055213	3.939 0.061564			1.710 0.062154
β _{Bathrooms} t	6.320	4.717				3.871
-	0.04914	0.052667				0.053714
β _{Notsquare} t	2.848	2.469				1.973
β _{SqFt}	0.007772	0.005708		0.007122		0.015597
t	18.951	6.457				4.733
βIndustrialLand	-0.00113					-0.000832
t	-1.663					-0.758
$\beta_{EmployDepriv}$	-0.02416	-0.02372			-0.0113	-0.016711
t	-5.899	-6.048			-4.836	-5.326
β1	0.00766	0.009199	0.001622	0.001981	0.006044	0.005016
t	1.352	4.447	5.059	1.577	1.221	1.587
β2	-0.00095	-0.00097	-0.00141	-0.00108	-0.00099	-0.000917
t	-3.148	-3.502	-2.349	-2.610	-3.756	-2.738
β3	0.000516	0.000485	0.001067	0.000606	0.000541	0.000498
t	1.953					1.745
β4	-3.79069	-3.87525	-3.9581	-4.07902	-3.85198	-3.91983
t	-23.445					-21.381
βPrimarySchool	0.000836		0.005957			0.003342
t	2.461	1.384				2.732
βsecondarySchool	0.588393	0.335556				-0.339676
<u>t</u>	6.215					-1.152
λ ₁	0.500048					0.419329
t	24.429	9.681	48.943	13.406	10.544	6.311

Appendix Table 2: Estimated parameters for Models I to VI, with t-statistics for each estimate.

Parameter	Model I	Model II	Model III	Model IV	Model V	Model VI
λ ₂	6.445736	6.508272	4.710182	6.026819	7.101111	4.938933
t	11.943	8.204	7.628	8 8.719	15.255	5.554
ξ	0.417822	0.425036	0.73667	0.737899	0.482328	0.531631
t	2.257	4.063	9.434	5.473	3.633	4.067
Ψ	-0.14056	-0.1213	-0.09503	-0.10093	-0.12659	-0.111524
t	-7.913	-2.794	-8.552	-3.568	-3.309	-2.156
β _{Periphery}					-0.60186	-0.002997
t					-1.925	-0.172
σ	0.075256	0.082311	0.095002	0.092192	0.079578	0.085398
	10.913	4.735	5 18.492	2 6.946	5.145	4.025
Log Likelihood	-2103.11	-2101.9	-2110.45	-2109.55	-2097.26	-2096.95
Ν	490	490	490	490	490	490

Variable		GCSE				Dist	Dist	Dist	Dist
	Price	Assigne d	GCSE Closest	Keystage2 Assigned	Keystage2 Closest	Assgn Pri	Close Pri	Assgn Sec.	Close Sec.
Price	1.000	0.138	0.182	0.202	0.215	0.089	0.142	-0.079	-0.146
GCSE Assigned	Secondary	1.000	0.435	0.450	0.409	0.163	0.203	0.264	0.168
GCSE Closest S	econdary		1.000	0.412	0.475	0.065	0.017	0.031	0.120
Keystage2 Assig	ned Primary			1.000	0.815	0.137	0.104	0.095	-0.067
Keystage2 Close	est Primary				1.000	0.104	0.071	-0.006	-0.109
Distance Assign	ed Primary					1.000	0.518	0.435	-0.069
Distance Closes	t Primary						1.000	0.285	-0.007
Distance Assign	ed Secondary	,						1.000	0.544
Distance Closes	t Secondary								1.000

Appendix Table 3: Correlations between school quality variables